# The EM Algorithm and its Application to Context Estimation Hauptseminar Machine Learning for Context Aware Computing

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Parametric techniques are often not sufficient.

- *Idea:* Mix several parametric distributions.
- Parameters of mixing components could be found easily if responsible component for every data point would be known.
- Next Idea: Treat this relationship as complementary data.
- Now the Expectation Maximization algorithm can iteratively yield good solutions.
- An Application to skin color blob tracking is shown.



#### Motivation

- Mixture Models
- The EM Algorithm: Overview
- The EM Algorithm: Expectation Step
- The EM Algorithm: Maximization Step
- Application: Skin Color Blob Tracking



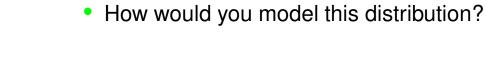
- General problem: Given a data set  $\mathcal{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$ , find a suitable distribution modeling the underlying random process  $p(\mathbf{x})$ .
- We have discussed *parametric models*  $p(\mathbf{x}|\theta)$  in detail, especially Gaussian models.
- We will see the power of *nonparametric models* (i.e. histograms, kernel-based methods) in later talks.
- Both techniques have advantages and drawbacks.
- Why not combine them?

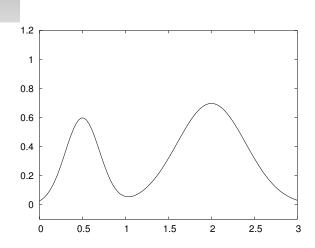


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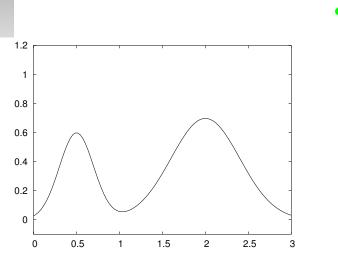
### Mixture Models







### Mixture Models

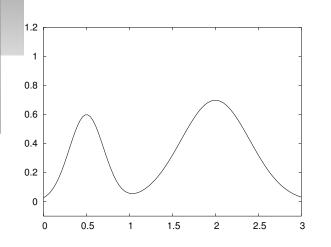


- How would you model this distribution?
- Mixture of two Gaussians:

$$p(x) = 0.3\mathbf{n}(\mu = 0.5, \sigma = 0.2) + 0.7\mathbf{n}(2.0, 0.4)$$



### **Mixture Models**



- How would you model this distribution?
- Mixture of two Gaussians:

$$p(x) = 0.3\mathbf{n}(\mu = 0.5, \sigma = 0.2) + 0.7\mathbf{n}(2.0, 0.4)$$

• General form of mixture model:

$$p(\mathbf{x}) = \sum_{j=1}^{M} P(j) p(\mathbf{x}|j)$$

with

$$\sum_{j=1}^{M} P(j) = 1$$
$$0 \le P(j) \le 1$$
$$\int p(\mathbf{x}|j) d\mathbf{x} = 1$$



• Log-likelihood for mixture density:

$$E = \ln(\mathcal{L}(\Theta|\mathcal{X})) = \ln\prod_{n=1}^{N} p(\mathbf{x}_n|\Theta) = \sum_{n=1}^{N} \ln\left(\sum_{j=1}^{M} P(j)p(\mathbf{x}_n|j)\right)$$

- Problem: sum of logarithms of sums is hard to optimize!
- We postulate existence of *unobserved data items*  $\mathcal{Z} = \{z_1, \ldots, z_N\}$  that tell us from which distribution every  $\mathbf{x}_n$  was taken.
- Now the likelihood is easier to optimize:

$$E = \ln(\mathcal{L}(\Theta|\mathcal{X}, \mathcal{Y})) = \ln(P(\mathcal{X}, \mathcal{Z}|\Theta)) = \sum_{n=1}^{N} \ln(P(\mathbf{x}_n|z_n)P(z_n))$$

$$= \sum_{n=1}^{N} \ln(P(z_n)p_{z_n}(\mathbf{x}_n|\theta_{z_n}))$$



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#### **Expectation step:**

- Guess some initialization values for parameters  $\theta_j$  of mixture models.
- Use  $\theta_j$ 's to calculate probability distribution of  $\mathcal{Z}$  depending on  $\mathbf{x}_n$ .
- Compute *expectation*  $\mathcal{E}$  of log likelihood  $E(\mathcal{X}, \mathcal{Z})$ .

#### Maximization step:

• Maximize  $\mathcal{E}(E(\mathcal{X}, \mathcal{Z}))$  with regard to the following parameters:

Distribution of  $\mathcal{Z}$ :  $P^{new}(j)$ 

Parameters of mixture model :  $\theta_j$ 

Iterate both steps:

Until a convergence criteria has been reached.



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 Start with computation of probability distribution for mixture weights:

$$P^{old}(j|\mathbf{x}_n) = \frac{P^{old}(j)p(\mathbf{x}_n|\theta_j^{old})}{\sum_{l=1}^M P^{old}(l)p(\mathbf{x}_n|\theta_l^{old})}$$

• Compute expectation by trying out all possibilities of  $\mathcal{Z}$ :

$$\mathcal{E}(E(\mathcal{X})) = \sum_{z_1=1}^{M} \sum_{z_2=1}^{M} \dots \sum_{z_N=1}^{M} E(\mathcal{X}, \mathcal{Z}) \prod_{n=1}^{N} P^{old}(z_n | \mathbf{x}_n)$$



### EM Algorithm: Expectation Step (2)

• Substitute

$$E(\mathcal{X}, \mathcal{Z}) = \sum_{n=1}^{N} \ln \left( P^{new}(z_n) p(\mathbf{x}_n | z_n) \right)$$

• Resulting in

$$\mathcal{E}(E(\mathcal{X}, \mathcal{Z})) = \sum_{n=1}^{N} \sum_{j=1}^{M} P^{old}(j | \mathbf{x}_n) \cdot \ln (P^{new}(j)) + \sum_{n=1}^{N} \sum_{j=1}^{M} P^{old}(j | \mathbf{x}_n) \cdot \ln (p^{new}(\mathbf{x}_n | \theta_j^{new}))$$

• Both terms are independent of each other and can be maximized individually.



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# EM Algorithm: Maximization Step (1)

Start with the maximization  $\mathcal{E}$  with regard to  $P^{new}(j)$ :

- Constraint  $\sum_{j=1}^{M} P(j) = 1$  handled by Lagrange multiplier  $\lambda$ .
- Set partial derivation to zero:

$$\frac{\partial}{\partial P^{new}(j)} \left[ \sum_{n=1}^{N} \sum_{j=1}^{M} P^{old}(j | \mathbf{x}_n) \cdot \ln\left(P^{new}(j)\right) + \lambda\left(\sum_{j=1}^{M} P^{new}(j) - 1\right) \right] = 0$$

• Resulting in  $\lambda = -N$  and

$$P^{new}(j) = \frac{1}{N} \cdot \sum_{n=1}^{N} P^{old}(j|\mathbf{x}_n)$$

Note: this result is independent of the underlying distribution



Now we try to maximize  $\mathcal{E}$  with regard to  $\theta_j^{new}$ .

- In general, this is extremely hard.
- We assume a very simplistic model: a *d*-dimensional Gaussian with covariance matrix  $\Sigma = \sigma^2 \cdot \mathbf{I}$ . Now  $\theta_j^{new} = (\mu_j, \sigma_j)$  and

$$p^{new}(\mathbf{x}_n | \theta_j^{new}) = \frac{1}{\left(2\pi\sigma_j^2\right)^{d/2}} \exp\left(-\frac{||\mathbf{x}_n - \mu_j||^2}{2\sigma_j^2}\right)$$

• Again, we have the Lagrange multiplier  $\lambda = -N$ , resulting in:

$$\mu_j^{new} = \frac{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)\mathbf{x}_n}{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)}$$
$$\sigma_j^{new2} = \frac{1}{d} \cdot \frac{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n) ||\mathbf{x}_n - \mu_j^{new}||^2}{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)}$$



### EM Algorithm: Maximization Step (3)

Computation of maximization is also possible with full covariance matrix:

$$P^{old}(j|\mathbf{x}_n) = \frac{P^{old}(j)\frac{1}{(2\pi)^{d/2}|\Sigma_j|^{1/2}} \cdot \exp\left(-\frac{1}{2}(\mathbf{x}_n - \mu_j^{old})^t \Sigma_j^{-1}(\mathbf{x}_n - \mu_j^{old})\right)}{\sum_{l=1}^M P^{old}(l)\frac{1}{(2\pi)^{d/2}|\Sigma_l|^{1/2}} \cdot \exp\left(-\frac{1}{2}(\mathbf{x}_n - \mu_l^{old})^t \Sigma_j^{-1}(\mathbf{x}_n - \mu_l^{old})\right)}$$

$$\mu_j^{new} = \frac{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)\mathbf{x}_n}{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)}$$

$$\Sigma_j^{new} = \frac{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)(\mathbf{x}_n - \mu_j^{new})(\mathbf{x}_n - \mu_j^{new})^t}{\sum_{n=1}^N P^{old}(j|\mathbf{x}_n)}$$



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Idea: Every pixel is composed of three parameters, *Red, Green* and *Blue*. For skin color, we could model the distribution probability individually in each of these parameters

#### Algorithm:

- 1. Set up mixture model consisting of M = 3 3D Gaussians.
- 2. Label some sample images by hand (click on skin colored pixels).
- 3. Let the EM algorithm determine the weights of the mixture model and the parameters  $(\mu_j, \sigma_j)$  of the Gaussian distributions.
- 4. Pre-compute a lookup table containing the probability of being skin color for every color in RGB space.
- 5. Online classification is now easy: do a lookup and classify current pixel if its probability is above a threshold.



# Application: Skin color blob tracking



#### Algorithm:

- 1. Extract skin colored pixels from image.
- 2. Set up Gaussian mixture model with three 2D Gaussians.
- 3. Feed the skin pixel's locations (x, y) in the EM algorithm.

**Result:** Three 2D Gaussians describing head and hands of a person on the image.



### Thank you

#### ...any Questions? ...

